

1 **Relationship between vitamin D status in the first trimester of the pregnancy**
2 **and gestational weight gain: A mediation analysis**

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5 **Running title:** Maternal vitamin D and gestational weight gain

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22

23 **Abstract**

24 **Objective**

25 To evaluate the total, and direct effects of vitamin D, measured by circulating 25-
26 hydroxyvitamin D [25(OH)D] levels, on GWG after adjustment for confounding variables, and
27 then assess the indirect effects by demonstrating the role of gestational age at birth as a mediator
28 in this association.

29 **Design**

30 A secondary analysis of data collected in a screening program in pregnancy.

31 **Setting and population**

32 Data collected in “Khuzestan Vitamin D Deficiency Screening Program in Pregnancy” was used
33 for the present study; it was included the data of 900 pregnant women referred to the health
34 centers of Shushtar (Khuzestan Province, Iran), whose vitamin D status during the third trimester
35 of pregnancy was available.

36 **Methods**

37 A mediation analysis was applied to detect the causal relationship between serum level of
38 25(OH)D, covariates (maternal age, parity, and baseline maternal weight), mediator (gestational
39 age), and outcome (GWG).

40 **Main outcome measures**

41 The main outcome measure of the study was gestational weight gain.

42 **Results**

43 The adjusted total effect of vitamin D on GWG was estimated 0.0699 (95%CI: 0.0537, 0.0849;
44 P=0.000). Although, an adjusted direct effect of vitamin D on GWG was not statistically

45 significant, the adjusted indirect effect of this micronutrient on GWG by considering gestational
46 age as a mediator was found to be significant [0.059 (95%CI: 0.048, 0.0708; P=0.000)]. Women
47 with severe vitamin D deficiency had the lowest speed as compared to moderate and normal
48 levels.

49 **Conclusion**

50 This study shows that maternal vitamin D status affects the gestational weight gain by reducing
51 the risk of preterm delivery.

52 **Keywords:** Vitamin D, gestational weight gain, preterm delivery

53 **Tweetable abstract**

54 The maternal vitamin D status can affect the gestational weight gain by reducing the risk of
55 preterm delivery.

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59 **Introduction**

60 Vitamin D, a fat-soluble vitamin and a pro-hormone, which can be synthesized from a steroid
61 precursor other or a dietary source, is essential for calcium and phosphorus homeostasis, and
62 bone mineralization.^{1,2} In addition to vitamin D function to maintaining bone health, this
63 micronutrient plays a crucial role in normal metabolism, and cellular growth, puberty,
64 reproduction, immune system regulation, and prevention of some medical conditions, such as
65 infectious, and cardio-metabolic disorders, cancers, depression, and cognitive deficits.¹⁻⁶
66 Vitamin D insufficiency measured by circulating 25-hydroxyvitamin D [25(OH)D] levels, the
67 best biomarker of vitamin D status, is considered a frequent disturbance among young women

68 particularly during pregnancy even in sunny regions, such as Iran, despite the intense sunlight in
69 these countries.⁷⁻¹⁰ The prevalence of maternal vitamin D deficiency is estimated to be 54%
70 worldwide.¹¹

71 Sufficient 25(OH)D concentrations are vital during pregnancy due to the increasing demand for
72 calcium during the growth and development of the fetus.^{12,13} Several studies have shown that
73 vitamin D deficiency in pregnant women is associated with adverse pregnancy outcomes, such as
74 preterm delivery, recurrent pregnancy losses, gestational diabetes¹⁴⁻¹⁶, preeclampsia¹⁷⁻¹⁹, primary
75 cesarean section rate^{20,21}, depression^{22,23}, and small-for-gestational-age (SGA) infants.^{12,24,25}

76 Moreover, a limited number of studies have suggested that 25(OH)D concentration can be
77 related to gestational weight gain (GWG), although the results of these studies are still debated
78 and conflicting²⁶⁻²⁸. For example, a recent study showed that the associations between circulating
79 levels of 25(OH)D and GWG only among pre-gestational overweight women²⁶, whereas other
80 studies revealed no relationship between maternal vitamin D status and GWG.^{27,28}

81 Generally, mechanisms involved in this association are complex and discussed; however, based
82 on the available documents, 25(OH) D influences GWG may influence through both direct and
83 indirect effects.²⁹⁻³⁴ Earlier studies have suggested that the effects of 25(OH) D on GWG may be
84 explained by biologic activities of this micronutrient on adipose tissue. Indeed, vitamin D
85 receptors (VDR) are present on human adipocytes and 25(OH) D appears to influence
86 lipogenesis, lipolysis, adipogenesis, and reducing adipose tissue inflammation.²⁹⁻³² Also due to
87 the anabolic effect of 25(OH)D on growth, vitamin D deficiency can be associated with impaired
88 maternal weight gain and fetal growth among vitamin D deficient mothers.³³ On the other hand,
89 it has shown that insufficient GWG is associated with an increased risk of adverse pregnancy
90 outcomes, such as preterm delivery.³⁵⁻³⁷ There is strong evidence demonstrating an increased risk

91 of preterm delivery in mothers with vitamin D deficiency^{34,38} and dilution of this risk in women
92 treated with supplementation³⁹; hence, preterm delivery due to insufficient serum level of
93 25(OH) D may mediate the association between this micronutrient and GWG. According to this
94 hypothesis, we aimed to evaluate the total, and direct effects of 25(OH)D level on GWG after
95 adjusting for confounding variables, and then assess the indirect effects by demonstrating the
96 role of gestational age at birth as a mediator in this association.

97 **Materials and Methods**

98 **Study design and participants**

99 This study was carried out on the data collected from the Khuzestan Vitamin D Deficiency
100 Screening Program in Pregnancy. The details of the study procedure have been reported
101 previously.³⁹ Briefly, this study was a stratified randomized field trial, consisting of two phases.
102 In the first phase, 1600 and 900 first-trimester pregnant women, referred to the health centers of
103 Masjed-Soleyman and Shushtar (Khuzestan Province, Iran), were recruited, respectively, and
104 fasting blood samples were collected. The serum samples of the participants in Shushtar were
105 stored and kept frozen at -80°C until further assays at the end of the study, whereas the vitamin
106 D status of participants in Masjed-Soleyman was immediately determined. In the second phase
107 of this study, the subjects with vitamin D deficiency from Masjed-Soleyman were assigned a
108 treatment regimen and received vitamin D3 supplementation until delivery. Other samples were
109 collected in the third trimester of pregnancy from all participants. Since participants from
110 Masjed-Soleman were treated with vitamin D supplementation, this study was conducted only on
111 participants referred to the health centers of Shushtar.
112 Participants received standard prenatal care, and both maternal and neonatal outcomes were
113 recorded. The adverse pregnancy outcomes included preterm delivery (birth at <37 weeks of

114 gestation), PE (systolic blood pressure >140 mmHg or diastolic blood pressure \geq 90 mmHg and
115 24-hour proteinuria \geq 0.3 g, initiated at >20 weeks of gestation), and GDM (glucose intolerance
116 first detected during pregnancy, based on the criteria of the International Association of Diabetes
117 and Pregnancy Study Groups). The study participants were classified into three groups according
118 to their serum concentration of 25(OH)D as severely deficient (<10 ng/mL), moderately deficient
119 (10 to 20 ng/mL), and >20 ng/mL.

120 **Clinical and laboratory measurements**

121 Trained examiners assessed the clinical and anthropometric measurements for all participants at
122 baseline and third trimester of pregnancy. All participants were interviewed for
123 sociodemographic, their history of pregnancies, and to obtain medical, obstetrics, and family
124 histories using pretested questionnaires. Adverse pregnancy outcomes were defined based on the
125 standard diagnostic criteria. At the time of data collection, women were asked about their history
126 of preeclampsia, based on a self-reporting questionnaire at each follow-up, details of which have
127 been previously published.³⁹

128 Serum levels of 25(OH)D were assayed for all participants at baseline and third trimester of
129 pregnancy. Circulating 25(OH)D levels were assayed using the ELISA method and a kit of
130 Immunodiagnosics Systems by Auto Analyzer (Human Corporation, Germany). The inter-assay
131 and intra-assay coefficients of variation were 3.891% and 3.37%, respectively (sensitivity of 5
132 nmol/L). Calibration of the instruments was done as per the manufacturer's instructions, and
133 validation studies were done before the test. Samples were analyzed by a single technician using
134 the same equipment throughout the study in a reference laboratory and were measured according
135 to standard operating procedures.

136 **Outcome of interest**

137 The gestational weight gain was considered as the outcome of interest of the study, which was
138 assessed at the end of each trimester of the pregnancy.

139 **Statistical analysis**

140 A mediation analysis was applied to detect the casual relationship between covariates, mediator,
141 and outcome variables. In this mediation analysis, gestational weight gain was considered as the
142 outcome of interest, vitamin D in the first trimester of pregnancy as the main exposure,
143 gestational age as a mediator, and maternal age, parity, and baseline maternal weight as the
144 potential confounding variables.

145 Once the crude analysis was applied to estimate the total, direct and indirect effects. The
146 estimated regression coefficient (95%CI) for total effect showed the overall mean of gestational
147 weight gain per increase of one unit vitamin D, regardless of any other covariates as the
148 mediators. The estimated regression coefficient (95%CI) for direct effect showed the mean of
149 weight gain during pregnancy per increase of one unit vitamin D, considering gestational age as
150 the confounding variable. Also, the estimated regression coefficient (95%CI) for indirect effect
151 showed the mean of weight gain during pregnancy per increase of one unit vitamin D,
152 considering gestational age as a mediator. The analysis was repeated to adjust maternal age,
153 parity, and baseline weight as the potential confounding variables. Bootstrap confidence intervals
154 and standard error were estimated through the bootstrap approach with replacement sampling of
155 10000. A direct acyclic graph (DAG) as the casual diagram was drawn to show the casual
156 relationships.

157 A trajectory plot with a fitted regression model was also used to show the trend of weight during
158 pregnancy trimesters for women with three group levels of vitamin D in the first trimester of
159 pregnancy (normal, moderate, and severe deficiency).

160 The analyses were conducted by SPSS software version 21 (SPSS Inc., Chicago, IL); PROCESS
161 v3.5 for the SPSS package was used to estimate total, direct, and indirect effects.

162 **Results** 163

164 Table 1 present the baseline characteristics and pregnancy outcomes of the study population. The
165 study participants had a mean age \pm SD of 29 ± 5 years and mean weight at baseline (SD) of 66.2
166 ± 7.1 kg. The median and IQR of the vitamin D level at baseline were 11.3 (8, 16.5) ng/ml. Table
167 2 shows the results of the crude mediation model. The total effect of vitamin D on GWG was
168 estimated 0.070 (95% CI: 0.055, 0.086; P= 0.000), which means the overall mean of gestational
169 weight gain was increased by 0.07 gram per each one-unit increase of vitamin D. After
170 adjustment for gestational age as confounding variable, vitamin D had no significant direct effect
171 on GWG [0.009 (95%CI: -0.007, 0.027; P=0.256)]. By considering the gestational age as a
172 mediator and estimating the indirect relationships through casual paths vitamin D \rightarrow GA \rightarrow WG,
173 vitamin D showed a significant effect on WG by 0.06 (95%CI: 0.049, 0.072; P=0.000); and the
174 mean of gestational weight gain was increased by 0.06 gram per each one-unit increase of
175 vitamin D. The causal relationships vitamin D in the first trimester of pregnancy and GWG is
176 presented in Figure 1.

177 Results of the mediation analysis adjusted for maternal age, parity, and baseline maternal weight
178 were presented in table 3. The adjusted total effect of vitamin D on WG was estimated 0.0699
179 (95%CI: 0.0537, 0.0849; P=0.000). Although, adjusted direct effect of vitamin D on WG was not
180 statistically significant [0.0105(95%CI: -0.007, 0.028; P=0.236)], the adjusted indirect effect of
181 this pro-hormone on WG considering GA as mediator was found to be significant [0.059(95%CI:
182 0.048, 0.0708; P=0.000)]. Figure 2 shows the causal relationships for this adjusted mediation
183 analysis.

184 Figure 3 shows the trajectory plot with the fitted regression model; it illustrates an increasing
185 trend of weight gain during pregnancy trimesters for women with different levels of 25(OH)D;
186 women with severe vitamin D deficiency had the lowest increasing speed as compared to
187 moderate and normal levels.

188 **Discussion**

189 This study was conducted to demonstrate the causal pathways between maternal vitamin D status
190 and gestational weight gain. Our findings demonstrate that 25(OH)D level in the first trimester of
191 pregnancy had a significant positive total effect on GWG, finding that remained significant after
192 adjustment for confounders, such as maternal age, parity, and baseline weight. While we found
193 no direct effect between 25(OH)D and GWG after adjustment for gestational age at birth (with or
194 without adjusting other confounders), when the gestational age was considered as a mediator, a
195 significant indirect effect was detected.

196 Overall, our study results showed the positive total effect of 25(OH)D on the GWG. We also
197 observed an increasing trend of GWG during pregnancy trimesters in both groups of women with
198 normal 25(OH)D level, and those with insufficiency, although mothers with severe vitamin D
199 deficiency had a lower GWG, compared to women with moderate deficiency and normal levels.
200 Our findings are in line with previous studies. A randomized controlled trial (RCT) conducted by
201 Hashemipour et al. showed that treatment of pregnant women with vitamin D deficiency resulted
202 in greater maternal weight gain during pregnancy. Another study conducted by Brooke et al.
203 showed that women treated with vitamin D supplementation gained weight faster in the late
204 trimester than those in the control group.⁴⁰ A meta-analysis of randomized controlled trials and
205 observational studies conducted by Thorne-Lyman and Fawzi²⁵ showed a greater average daily
206 weight gain in the third trimester of gestation among women supplemented with vitamin D. In

207 contrast to our results, a cohort study conducted by Figueiredo et al. showed that women who
208 had vitamin D inadequacy presented a higher increase in total gestational weight gain compared
209 to those with vitamin D adequacy and this association was present only in overweight women ²⁶;
210 they explained their results by the fact that vitamin D modulates adipogenesis and apoptosis and
211 thus regulates adipose tissue growth and also inflammation. ^{41,42} A possible explanation for the
212 difference between the results of our study and those of Figueiredo et al. is that in this study we
213 measured vitamin D once at 14 gestational weeks, while Figueiredo et al. measured it three times
214 throughout pregnancy and also the vitamin D inadequacy group in their study had a small sample
215 size. Also, in contrast with our results, Shakeri et al. ²⁷ and Nobles et al. ²⁸ found no relationship
216 between maternal vitamin D status and gestational weight gain. This discrepancy could be due to
217 differences in studies designs. For example, Shakeri et al. ²⁷ measured vitamin D levels in the
218 third trimester, while our study measured it at the first trimester of the pregnancy. Also, the study
219 of Nobles et al. ²⁸ conducted on women at risk for gestational diabetes mellitus, whereas this
220 study has conducted on healthy women. None of these studies adjusted the results for
221 confounding factors, while we adjusted our results by maternal age, parity, and baseline weight.
222 Although some previous studies have been reported overall effects of 25(OH)D level on GWG,
223 the exact mechanisms through which vitamin D may affect GWG have not yet clarified. ²⁶⁻²⁸ In
224 other words, it has not understood whether this prohormone has a direct, indirect, or both effects
225 on GWG. It is well documented that adipose tissue, the main storage site for vitamin D,
226 expresses vitamin D receptors (VDR) and enzymes involved in vitamin D metabolism. ⁴³⁻⁴⁵
227 Vitamin D has both stimulating and inhibiting effect on adipogenesis and modulatory effect on
228 adipose tissue inflammation and energy homeostasis. ⁴⁶ Moreover, vitamin D deficiency is
229 associated with anorexia and malaise which may explain the poor weight gain among pregnant

230 women with vitamin D inadequacy.⁴⁷ Also due to the anabolic effect of vitamin D on growth,
231 vitamin D deficiency is associated with impaired maternal weight gain and fetal growth among
232 vitamin D deficient mothers.³³ Despite the mentioned mechanisms, this study showed no direct
233 effect of 25(OH)D on GWG. Since several factors could affect the GWG, we adjusted our results
234 for important potential confounders, such as maternal age, parity, and baseline weight; our
235 results remained unchanged after these adjustments.

236 Interestingly, when gestation age at birth was considered as a mediator factor, our results
237 detected an indirect association between 25(OH)D level and GWG. Indeed, the role of
238 gestational age as a mediator can be explained by the possible effect of vitamin D on decreasing
239 the risk of preterm delivery. It has been shown that vitamin D supplementation could decrease
240 the risk of preterm delivery up to 40%.³⁹ Other reports also found an inverse association between
241 maternal vitamin D and preterm delivery.^{34,38} Vitamin D is a potent regulator of inflammation in
242 the placenta, this may explain the linkage between vitamin D with pathologic conditions such as
243 pre-eclampsia⁴⁸. Vitamin D also regulates target genes associated with proper implantation of
244 the placenta and is important for pregnancy maintenance through being related to calcium
245 metabolism in myometrium.^{49,50} In addition, vitamin D influences other aspects of immunity,
246 especially the stimulation of antimicrobial innate immune response^{51,52} and direct role in the
247 production of antimicrobial peptides such as cathelicidin⁵³ thus play an important role in
248 preventing infection during pregnancy. Intrauterine infections lead to preterm delivery through
249 mechanisms related to activation of the innate immune system.⁵⁴ The mechanism through which
250 vitamin D reduces the risk of preterm delivery could be explained due to the effect of vitamin D
251 on innate immune and antimicrobial responses in placental cells and preventing infections during

252 pregnancy. These mechanisms emphasize on the indirect effects of vitamin D on the GWG,
253 which are mediated by gestational age at birth.
254 To the best of our knowledge, this is the first population-based study to determine the total,
255 direct, and indirect effects of serum level of 25(OH)D and GWG. The main strengths of this
256 study were the population-based design, the relatively large sample size, and the use of
257 appropriate statistical methods for data analysis, with adjustments for important confounders
258 such as age, parity, and maternal weight at baseline. It is important to emphasize that the present
259 study has some limitations. First, as data collection was carried out throughout the year, we were
260 unable to adjust for seasonal variance; second, we were not able to recruit liquid chromatography
261 technique to quantify 25(OH)D values; however, the ELISA technique is considered as a reliable
262 method when performed by experienced staff.⁵⁵ Finally, due to the diversity of food supplies, we
263 had no information on vitamin D dietary intakes in our participants.

264 **Conclusion**

265 In conclusion, this study shows that maternal vitamin D status affects the gestational weight gain
266 by reducing the risk of preterm delivery. Therefore, the detection and treatment of women with
267 vitamin D inadequacy can improve the trend of their weight gain by reducing the risk of preterm
268 delivery. However, further studies with a prospective design and more comprehensive measures
269 are warranted to disentangle the association between vitamin D status and total gestational
270 weight gain.

271 **Disclosure of interests**

272 The authors declare no competing interests. Completed dis-closure of interest forms are available
273 to view online assupporting information.

274 **Contribution to authorship**

275 MA was involved in the study design, managed the literature search, interpretation of data, and
276 manuscript drafting. MR was involved in the study design and data collection, carried out the
277 sample analysis, and manuscript drafting. RBY was contributed in statistical analyses,
278 interpreting of data, and manuscript drafting. AF was involved in searching literature,
279 interpretation of data, and manuscript drafting. MS was contributed in the study design,
280 interpretation of data, and manuscript drafting. FRT was involved in the study conception and
281 design and carried out the analysis and interpretation of data, managed the literature search, and
282 manuscript drafting.

283 **Details of ethics approval**

284 This study was approved by the Ethics Committee of the Research Institute of Endocrine
285 Sciences (IR.SBMU.ENDOCRINE.REC.1399.005) and a written informed consent was obtained
286 from all participants.

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450 **Table 1.** The baseline characteristics and pregnancy outcomes of the study population.
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Quantitative variables	Mean (SD) or Median (IQR)
Age	29 (5)
Maternal weight at baseline (kg)	66.2 (7.1)
Maternal weight at the second trimester of the pregnancy (kg)	70.9 (7.2)
Maternal weight at the third trimester of the pregnancy (kg)	73.4 (7.4)
Gravidity	2 (1, 3)
Parity	1 (0, 2)

Vitamin D level at baseline (ng/mL)	11.3 (8, 16.5)
Vitamin D level at delivery (ng/mL)	11 (8, 17)
Neonatal vitamin D level at birth (ng/mL)	10 (6.4, 15.2)
Categorized variables	n (%)
Education status	
Illiterate	17 (1.9)
Under diploma	671 (76.5)
Diploma or academic	189 (21.6)
Occupational status	
Household	629 (71.7)
Self-employed	113 (12.9)
Employed	135 (15.4)
Stillbirth	
No	870 (99.2)
Yes	7 (0.8)
Preeclampsia	
No	739 (84.3)
Yes	138 (15.7)
Gestational diabetes	
No	823 (93.8)
Yes	54 (6)
Preterm delivery	
No	752 (85.7)
Yes	125 (14.3)

Abbreviations: SD: standard deviation; IQR: interquartile range; n: number

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Table 2. Crude mediation model to estimate the total, direct and indirect effects of Vitamin D on WG and GA as a mediator.

Parameter estimated	Beta	*Boot Std. Error	*95% Wald Confidence Interval		
			Boot Lower	Boot Upper	Sig.
Total effect of Vitamin D on WG	0.070	0.0080	0.055	0.086	0.000*
Direct Effect of vitamin D by adjusting Gestational age as confounding variable	0.009	0.0087	-0.007	0.027	0.256
Indirect effect of vitamin D by considering Gestational age as a Mediator	0.06				0.000*

0.0059 0.0492 0.0722

472 * Number of bootstrap samples for percentile bootstrap confidence intervals and Standard Error: 10000
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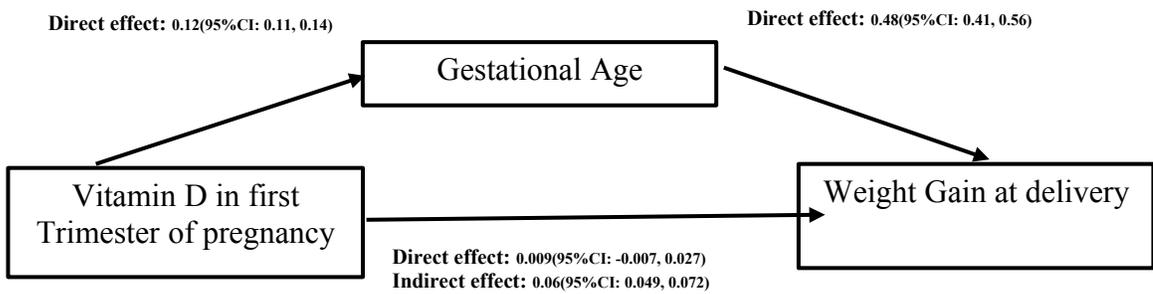
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494 **Table 3.** Adjusted the mediation model to estimate the total, the direct and indirect effect of Vitamin D on
 495 WG adjusting by maternal age, parity, and baseline weight as potential confounding and GA as a
 496 mediator.
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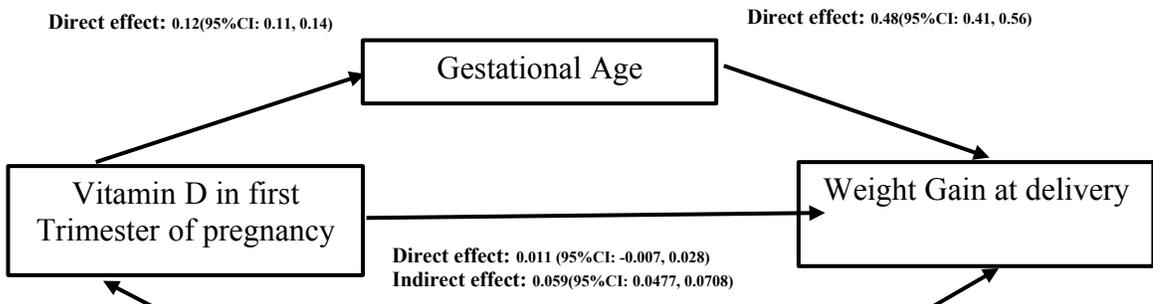
Parameter estimated	Beta	*Boot Std. Error	*95% Wald Confidence Interval		Sig.
			Boot Lower	Boot Upper	
The total effect of Vitamin D on WG adjusted by maternal age, parity, and baseline weight as confounding variables	.0699	.0079	.0537	.0849	0.000*

Direct Effect of vitamin D by adjusting Gestational age, maternal age, maternal age, parity, and baseline weight as confounding variables	.0105	.0087	-.0068	.0275	0.236
The effect of vitamin D by considering Gestational age as a Mediator, and maternal age, maternal age, parity, and baseline weight as confounding variables	.0585	.0059	.0477	.0708	0.000*

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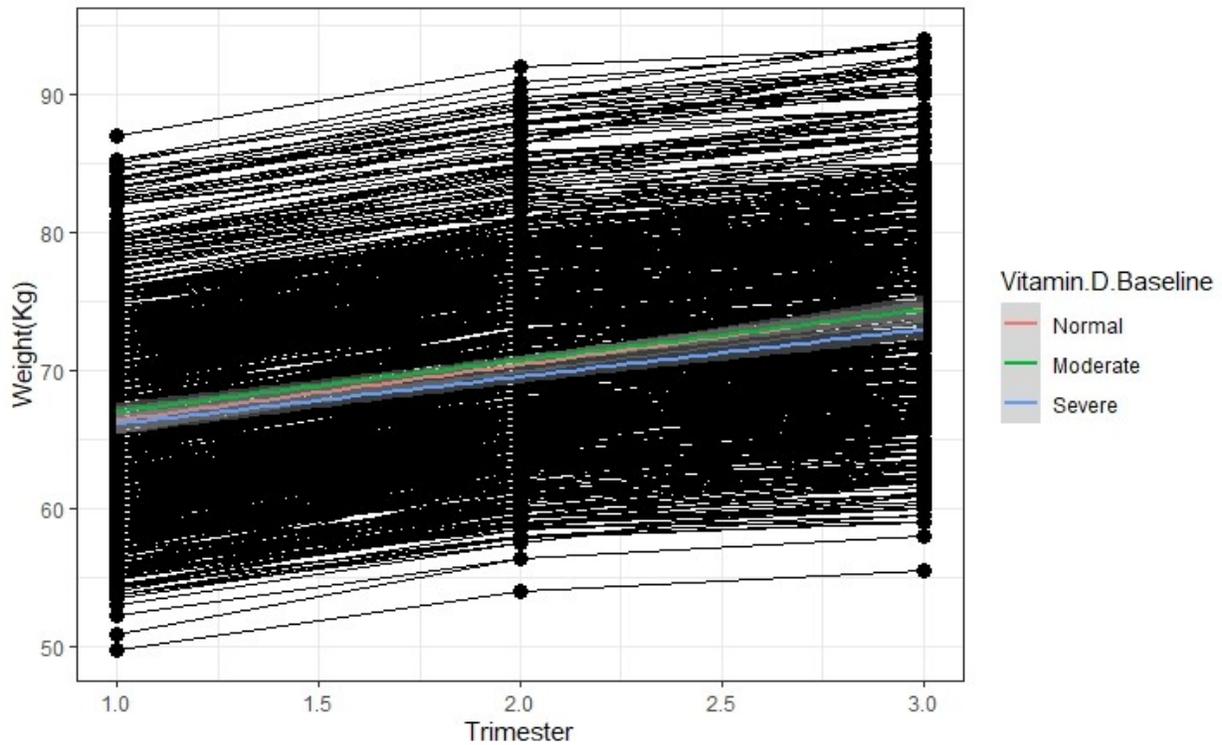
539 **Figure 1.** A crude mediation model with a single mediator variable Gestational age causally located between
540 Vitamin D in the first trimester of pregnancy and Weight gain during pregnancy.
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Maternal age, parity,
baseline weight

Figure 2. An adjusted mediation model with a single mediator variable Gestational age causally located between Vitamin D in the first trimester of pregnancy and Weight gain during pregnancy and Maternal age, Parity and baseline weight as the potential confounding variables.



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Figure 3. Trajectory plot with a fitted regression model to show the trend of weight in pregnancy trimesters for women with three groups of vitamin D at baseline (Normal, moderate, and severe deficiency).

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