

# **Estimating the CO<sub>2</sub> fertilization effect on extratropical forest productivity from Flux-tower observations**

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## **Key Points:**

- We present a novel statistical method to disentangle the variability of photosynthetic rates related to climate and non-climate drivers
- The analysis from 38 eddy covariance sites shows a  $16.4 \pm 4\%$  increase in photosynthetic carbon uptake for a 100 ppm rise in atmospheric CO<sub>2</sub>
- Our statistical method is successfully validated against idealized model simulations with and without increasing CO<sub>2</sub>

## Abstract

The land sink of anthropogenic carbon emissions, a crucial component of mitigating climate change, is primarily attributed to the CO<sub>2</sub> fertilization effect on global gross primary productivity (GPP). However, direct observational evidence of this effect remains scarce, hampered by challenges in disentangling the CO<sub>2</sub> fertilization effect from other long-term drivers, particularly climatic changes. Here, we introduce a novel statistical approach to separate the CO<sub>2</sub> fertilization effect on GPP and daily maximum net ecosystem production (NEP<sub>max</sub>) using eddy covariance records across 38 extratropical forest sites. We find the median stimulation rate of GPP and NEP<sub>max</sub> to be  $16.4 \pm 4\%$  and  $17.2 \pm 4\%$  per 100 ppm increase in atmospheric CO<sub>2</sub> across these sites, respectively. To validate the robustness of our findings, we test our statistical method using factorial simulations of an ensemble of process-based land surface models. We acknowledge that additional factors, including nitrogen deposition and land management, may impact plant productivity, potentially confounding the attribution to the CO<sub>2</sub> fertilization effect. Assuming these site-specific effects offset to some extent across sites as random factors, the estimated median value still reflects the strength of the CO<sub>2</sub> fertilization effect. However, disentanglement of these long-term effects, often inseparable by timescale, requires further causal research. Our study provides direct evidence that the photosynthetic stimulation is maintained under long-term CO<sub>2</sub> fertilization across multiple eddy covariance sites. Such observation-based quantification is key to constraining the long-standing uncertainties in the land carbon cycle under rising CO<sub>2</sub> concentrations.

## Plain Language Summary

Through photosynthesis, plants convert CO<sub>2</sub> and water into sugars and oxygen using solar energy, one of the most important chemical reactions on Earth. Human-made carbon emissions are increasing atmospheric CO<sub>2</sub> levels, impacting global photosynthesis. The additional carbon is believed to have a fertilizing effect on photosynthesis, causing vegetation to absorb a significant portion of the emitted CO<sub>2</sub>. However, the strength of this CO<sub>2</sub> fertilization effect on photosynthesis is uncertain, but is a crucial factor in determining the future trajectory of atmospheric CO<sub>2</sub> concentrations. In this study, we introduce a new statistical method to quantify the increase in photosynthetic carbon uptake, stimulated by rising CO<sub>2</sub>, based on measurements from 38 forest sites. Our results reveal that a 100 ppm increase in CO<sub>2</sub> enhances photosynthesis by approximately 16%. Validation of the statistical method with artificial model experiments

supports the robustness of our findings. Our study improves the understanding of the impacts of human-made CO<sub>2</sub> emissions on the global carbon cycle.

## 1 Introduction

Forests play a crucial role in the global carbon cycle, acting as a significant sink for anthropogenic carbon emissions. Approximately 25% of annual carbon emissions are estimated to be sequestered and stored by forests via photosynthesis, with boreal and temperate forests making substantial contributions (Pan et al., 2011). The physiological effects of increasing atmospheric carbon dioxide (CO<sub>2</sub>) on plant productivity, known as the CO<sub>2</sub> fertilization effect, are expected to stimulate photosynthesis and drive the enhanced carbon uptake. However, obtaining observational evidence for these effects in natural ecosystems and understanding how this process has changed historically remains a key knowledge gap.

Multiple lines of evidence support an enhancement in photosynthesis (or gross primary production; GPP) in response to an increase in CO<sub>2</sub>: CO<sub>2</sub> enrichment experiments (Norby et al., 2010; Walker et al., 2021), ecosystem monitoring (Keenan et al., 2013; Fernández-Martínez et al., 2017; Mastrotheodoros et al., 2017) and indirect proxies based on long-term atmospheric carbonyl sulfide records (Campbell et al., 2017) or isotopomer signal (Ehlers et al., 2015). Process-based land-surface models, which simulate the physiological responses of vegetation to environmental changes, also predict a stimulation of photosynthesis with increasing CO<sub>2</sub> levels (Sitch et al., 2015). However, multi-model projections of the CO<sub>2</sub> effect on long-term GPP diverge considerably due to uncertainties in process parameterizations and feedback mechanisms, particularly in response to meteorological extremes and climatic changes associated to rising CO<sub>2</sub> (Zaehle et al., 2005; De Kauwe et al., 2013; Rogers et al., 2014). Constraining the CO<sub>2</sub> fertilization effect in models through direct observational evidence is a long-called-for necessity to advance our understanding of carbon cycling and essential for more reliable future projections of carbon sequestration.

The global network of eddy covariance (EC) flux towers observing the exchange of CO<sub>2</sub> at the ecosystem scale provides a valuable resource, as it has accumulated sufficiently long time series data to potentially provide direct evidence of the CO<sub>2</sub> fertilization effect (Knauer et al., 2017; Baldocchi, 2020; Zhan et al., 2022). Previous studies have attempted to attribute the CO<sub>2</sub> fertilization effect on GPP by utilizing EC records. However, these studies have not adequately

accounted for confounded drivers of GPP, potentially leading to a misattribution, or only relied on an indirect use of measurements for attribution, such as using EC data to calibrate a predefined model. For instance, (Chen et al., 2022) used an eco-evolutionary optimality framework to reproduce EC-inferred GPP and subsequently attribute the CO<sub>2</sub> fertilization effect on GPP. Their analysis estimated a global GPP enhancement of 63 g C m<sup>-2</sup> yr<sup>-1</sup> from 2001-2014 due to rising CO<sub>2</sub>. Similarly, (Ueyama et al., 2020) utilized a model constrained with data from 104 EC towers and estimated a 12.4 g C m<sup>-2</sup> yr<sup>-1</sup> increase in GPP. (Fernández-Martínez et al., 2017), employing generalized mixed linear models, attributed an increase of  $11.2 \pm 2.5$  g C m<sup>-2</sup> yr<sup>-1</sup> in the GPP from 1995-2011 at 23 forest sites to CO<sub>2</sub>. These variations emphasize the importance of disentangling the CO<sub>2</sub> effect on GPP directly, *i.e., not using predefined model structures*, and by carefully considering confounding drivers in leveraging the continuously growing EC records.

In this study, we aim to disentangle the CO<sub>2</sub> fertilization effect on photosynthetic uptake directly and exclusively from long-term multi-site flux measurements and accompanying meteorological data. Several factors, such as CO<sub>2</sub>, climate changes, land-use and land-cover changes, affect ecosystem productivity and are correlated and confounded on the multi-decadal time-scale. We introduce the so-called *GPP residual method* that statistically captures the sensitivity of GPP to CO<sub>2</sub> and climate variables at different time scales to account for co-linearities among the drivers. First, we detrend the GPP time series to separate the long-term variability of GPP (the trend) primarily driven by CO<sub>2</sub> and climate, from the shorter-term variability (anomalies) primarily driven by climate variabilities. The method estimates the sensitivity of GPP to climate, referred to as the  $\gamma$  effect with temperature change as the proxy for climatic change, based on these anomalies. Next, we can quantify the  $\gamma$  effect based on long-term changes assuming the sensitivity remains consistent over the time scales of a few decades. The difference between the original observed GPP trend and the inferred GPP trend due to climate changes yields the unexplained GPP residual, which can be attributed to the long-term CO<sub>2</sub> effect on GPP, here referred to as the  $\beta$  effect. Specifically, we define the  $\beta$  factor as the relative change in GPP (%) per 100 ppm increase in CO<sub>2</sub>, and the  $\gamma$  factor as the relative change in GPP (%) per Kelvin increase in surface air temperature. It's important to note that the GPP residual may also include other long-term effects specific to individual sites, such as signals related to nitrogen deposition or land management. First, we validate the robustness of the *GPP residual method* using site-

level simulations with the QUINCY model (QUantifying Interactions between terrestrial Nutrient CYcles and the climate system; Thum et al., 2019). We use the *GPP residual method* to further estimate the  $\beta$  and  $\gamma$  effects in both GPP and daily maximum net ecosystem production ( $\text{NEP}_{\text{max}}$ ) using long-term EC records. At last, we compare the results from EC records and factorial simulations from a set of land-surface models (TRENDY version 9; Sitch et al., 2015). Additionally, we discuss the relevance of other potential long-term impacts on GPP in more detail.

## 2 Materials and Methods

### 2.1 Eddy covariance data

This study comprises 38 forested eddy covariance (EC) sites (Table S1), where  $\text{CO}_2$  flux data and meteorological data are collected by flux towers from Integrated Carbon Observation System (ICOS; Rebmann et al., 2018), and AmeriFlux (Novick et al., 2018). We focus on tree-dominated ecosystems as they exhibit less sensitivity to short-term climate variability compared to grass-dominated ecosystems. We focus on tree-dominated ecosystems due to the heightened sensitivity of grass-dominated ecosystems to short-term climate variability, adding complexity to the disentanglement of their response to  $\text{CO}_2$  (Hovenden et al., 2014; Reich et al., 2018). The sites span geographically across Europe and North America. The forest types can be broadly classified into: deciduous broadleaf forest (DBF; 12 sites), evergreen need-leaved forest (ENF; 20 sites), and mixed deciduous–coniferous forest (MF; 6 sites).

We obtain long-term recorded ( $\geq 10$  years) eddy covariance data at daily scale and net ecosystem production (NEE) at half-hourly scale from 1994 to 2022 (Table S2). GPP is estimated from the nighttime partitioning algorithm (Reichstein et al., 2005). Meteorological variables used in this study include temperature, incoming shortwave radiation and vapour pressure deficit (VPD). Due to the limited depth of soil moisture measurements (Yu et al., 2022), we calculate a water availability index (WAI; Tramontana et al., 2016) following a bucket model approach. The maximum cumulative water deficit (Aragão et al., 2007) represents the available water content (awc). WAI is calculated as the balance of precipitation recharge and observed evapotranspiration as follows:

$$input(t) = \min(Precipitation(t), awc - WAI(t)) \quad (1)$$

$$WAI(t + 1) = \max((WAI(t) + input(t) - Evapotranspiration(t)), 0) \quad (2)$$

Where  $t$  represents the timestep  $t$ . We exclusively include data with daily quality control indicators for NEE and meteorological variables surpassing 0.6, denoting a 60% or higher percentage of measured and high-quality gap-filled data. In summary, EC sites are selected for this study based on three criterias: (a) only sites dominated by tree-ecosystems are selected; (b) there has to be a long-term record of EC observations ( $\geq 10$  years) after the quality control when (c) at least 60% data per day is measured or gap-filled with good quality. In total we estimate  $\beta$  and  $\gamma$  for 228 site-months.

Maximum leaf area index ( $LAI_{max}$ ,  $m^2 m^{-2}$ ), forest age (years), site-specific disturbance or management information are collected from the literature (Musavi et al., 2017; Besnard et al., 2018; Flechard et al., 2020; Migliavacca et al., 2021), the BADM product, and/or site principal investigators.

Due to the systematic biases of the atmospheric  $CO_2$  measurements in the eddy covariance data, for consistency, we replace the  $CO_2$  measurement with the  $CO_2$  product CAMS CF-1.6 (Chevallier et al., 2005, 2010) from the nearest pixel to each EC site. The  $CO_2$  reanalysis data spans from 1994 to 2022 with daily resolution, thus sufficient to match the time period of the eddy covariance data.

## 2.2 Estimating $\beta$ and $\gamma$ using the *GPP residual method*

We develop the *GPP residual method* (Figure 1) to isolate the  $CO_2$  fertilization effect ( $\beta$ ) from other confounding factors (*e.g.*, *climate*).  $\beta$  is inferred for each site and each month-of-year separately, using the median values of GPP and hydro-meteorological data across five-day intervals within the considered months to filter out synoptic weather variability and its impact on GPP dynamics. The calculation of  $\beta$  consists of three steps: (1) Data preparation (Figure 1b). The growing season when plant photosynthesis is active is defined based on the mean seasonal cycle of GPP (averages by day-of-year) across the time series. A month is considered within the growing season, if there are more than twenty days when GPP is greater than 25% of the

maximum of GPP as inferred from the mean seasonal cycle. Within each month, the median values of all variables are retrieved for every five-day interval. We then calculate anomalies using the median values by subtracting the long-term trend of the linear-fit for each month-of-year (*e.g.*, *July in 1999, ..., 2020*). We rescale the anomalies of all variables by adding the average value across the considered time period. The rescaling allows the random forest model in step (2) is trained and applied at an identical magnitude to extrapolate; (2) Model training and predicting climatic effects on GPP (Figure 1c). We train a random forest regression model for GPP anomalies using anomalies of hydro-meteorological variables (*i.e.*, *temperature (Temp)*, *incoming shortwave radiation (SW<sub>in</sub>)*, *vapor pressure deficit (VPD)*, *water availability index (WAI)*, *growing degree days (GDD)*). We use the model to predict GPP using the actual hydro-meteorological data (including both trends and anomalies) at each month-of-year. The predicted GPP (GPP<sub>climatic</sub>) thus only reflects the effect of climate. Next to the random forest regression model, we call a multivariate linear regression model to test the robustness of the results from the random forest regression model (Figure S1); (3) Isolating non-climatic effects on GPP (Figure 1d). The non-climatic effects on GPP (GPP<sub>residual</sub>) are derived by removing the GPP<sub>climatic</sub> from the actual GPP time series (GPP<sub>residual</sub> = GPP - GPP<sub>climatic</sub>). GPP<sub>baseline</sub> is calculated as the mean GPP over the first two years in the considered time series. The relative change of GPP in response to CO<sub>2</sub> (*i.e.*,  $\beta$ ) is derived as the trend of the linear-fit between CO<sub>2</sub> concentration and GPP<sub>residual</sub> in relative to GPP<sub>baseline</sub>. Similarly, the sensitivity of GPP to temperature (*i.e.*,  $\gamma$ ) is derived as the trend of linear-fit between temperature and GPP<sub>climatic</sub> in relative to GPP<sub>baseline</sub>:

$$\beta = \frac{\frac{\Delta GPP_{residual}}{GPP_{baseline}}}{\Delta CO_2} \times 100\% \quad (3)$$

$$\gamma = \frac{\frac{\Delta GPP_{climatic}}{GPP_{baseline}}}{\Delta Temp} \times 100\% \quad (4)$$

The advantage of the *GPP residual method* is that we separate the confounding factors at different time scales, thus, it can solve the issue of multicollinearity to some extent, when the independent variables are highly correlated to one another. To show the different results yield from the *GPP residual method* and a multivariate regression method, we adopt a simple multivariate regression model as the following:

$$\frac{GPP}{GPP_{baseline}} * 100\% = \beta \times CO_2 + \gamma \times Temp + \rho_{VPD} \times VPD + \rho_{WAI} \times WAI + \rho_{GDD} \times GDD + \rho_{SW_{in}} \times SW_{in} \quad (5)$$

Where  $\beta$  is the sensitivity of GPP to  $CO_2$ , and  $\gamma$  is the sensitivity of GPP to temperature. After obtaining  $\beta$  values for each site-month, we calculate the median  $\beta$ . This approach helps to mitigate the influence of outliers. We further estimate the uncertainty of the median  $\beta$  using the bootstrap method. By repeatedly sampling from the considered  $\beta$  distribution, we create multiple bootstrap samples. Each sample is then used to calculate the median  $\beta$ . The standard deviation across these bootstrap estimates provides an estimate of the uncertainty associated with median  $\beta$ . We calculate the median  $\gamma$  and its uncertainty in the same way.

To consider the seasonal and spatial variation of GPP, we further calculate annual  $\beta$  and  $\gamma$  by aggregating monthly  $\beta$  and  $\gamma$  weighted by monthly  $GPP_{baseline}$ :

$$Annual \beta = \sum_{i=1}^n \beta_{month_i} \times \frac{GPP_{baseline_i}}{GPP_{gs}} \quad (6)$$

Where  $i$  represents a specific month, and  $n$  is the total number of months.  $GPP_{gs}$  is the sum of baseline GPP across the considered growing season. Similarly, the annual  $\beta$  and  $\gamma$  at each site can be further aggregated across space:

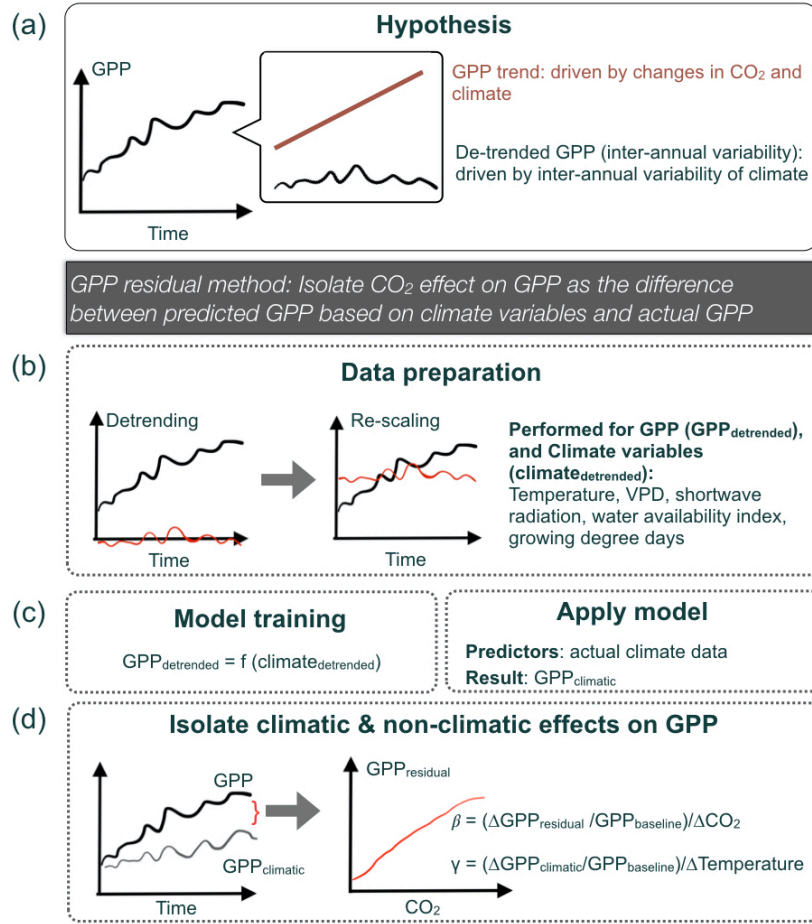
$$mean \beta = \sum_{i=1}^n Annual \beta_i \times \frac{GPP_{gs_i}}{GPP_{sum}} \quad (8)$$

$$mean \gamma = \sum_{i=1}^n Annual \gamma_i \times \frac{GPP_{gs_i}}{GPP_{sum}} \quad (9)$$

Where  $i$  represents a specific site, and  $n$  is the total number of sites.  $GPP_{sum}$  is the sum of baseline GPP across all sites. To assess the robustness of the median  $\beta$  or  $\gamma$  values and determine if they are influenced by site selection, we compare the mean  $\beta$  or  $\gamma$  calculated across all sites, weighted by baseline GPP, with the median  $\beta$  or  $\gamma$  derived from the distribution of monthly  $\beta$  or  $\gamma$  values. If the median value remains relatively stable and comparable to the mean value across all sites, it suggests that the selection of sites does not significantly impact the robustness of the



215 median  $\beta$  or  $\gamma$  estimations.



216

217 **Figure 1.** Schematic of the statistical *GPP residual method* to isolate the CO<sub>2</sub> fertilization effect  
 218 and climatic effect in observational data of GPP. (a) Hypothesis. The overall goal of the *GPP*  
 219 *residual method* is to isolate the CO<sub>2</sub> fertilization effect on GPP by removing long-term climate  
 220 effects inferred from short-term variability. (b) Data preparation. All the time series of climate  
 221 variables and GPP are detrended and individually rescaled to the long-term mean of each  
 222 variable. The black lines denote the actual time series for each variable, and the red line denotes  
 223 the detrended time series. (c) A random forest model or multivariate linear regression model is  
 224 trained to learn the sensitivity of GPP to the climate variables based on the detrended time series,  
 225 i.e., interannual variability. (d) The trained model predicts the long-term changes in GPP caused  
 226 by climate changes using the original time series of climate predictors, including the long-term  
 227 trend. The non-climate-induced effect on GPP is therefore estimated from the residual of  
 228 absolute GPP minus climate-induced GPP, shown as the red line.

229 The EC technique allows for direct measurement of NEE, which is the difference between GPP  
 230 and ecosystem respiration (RECO):

231 
$$NEE = RECO - GPP \quad (10)$$

The maximum net ecosystem production ( $NEP_{max}$ ) is defined as the negative sign of the minimum NEE during a day from half-hourly measurement:

$$NEP_{max} = -NEE_{min} \quad (11)$$

In addition to GPP, we further estimate  $\beta$  or  $\gamma$  for  $NEP_{max}$  following the same method.

### 2.3 Validate the *GPP residual method* with a land surface model

We use the terrestrial biosphere model QUINCY (QUantifying Interactions between terrestrial Nutrient CYcles and the climate system; Thum et al., 2019), which has been evaluated against a subset of FLUXNET sites across large geographical ranges and different ecosystem types, to validate the *GPP residual method*. We perform two simulations with identical climate but varying  $CO_2$  concentrations (transient  $CO_2$  as observed, and constant  $CO_2$  at levels of the year 1988) at 166 forested sites distributed across the globe. The model setup and model simulations are identical with the “freeze- $CO_2$  experiment” in Zhan et al. (2022). For better comparison with eddy covariance records, we take the last twenty years (1999-2018) in the simulations as the time period of the validation.

The advantage of this method is that we can compare the  $\beta$  estimated by the *GPP residual method* ( $\beta_{estimated}$ ) with  $\beta$  modeled by QUINCY ( $\beta_{QUINCY}$ ), which is regarded as a surrogate truth that represents theory of photosynthetic responses to  $CO_2$ , climate and water availability.  $\beta_{QUINCY}$  is calculated as the sensitivity of the  $CO_2$ -induced change in GPP to  $CO_2$  concentration, in which GPP is calculated as the difference between simulations forced with transient  $CO_2$  and constant  $CO_2$  during the considered time period.  $\beta$  is calculated for each site and each month-of-year. The selection of months follows the same rule as the data preparation in the previous section.  $\beta_{estimated}$  is calculated using the *GPP residual method* elaborated in the previous section from the simulation forced with transient  $CO_2$ .

We evaluate the agreement of  $\beta_{estimated}$  and  $\beta_{QUINCY}$  for each forested site (166 sites in total) in the model. We use the root-mean-square error (RMSE) to measure the difference between  $\beta_{estimated}$  and  $\beta_{QUINCY}$  across the growing season. The RMSE of  $\beta$  estimation per site is calculated as:

$$RMSE(\beta) = \sqrt{\frac{\sum_{i=1}^n (\beta_{estimated_i} - \beta_{QUINCY_i})^2}{n}} \quad (10)$$

where  $n$  is the number of months when  $\beta$  is estimated;  $\beta_{estimated_i}$  and  $\beta_{QUINCY_i}$  is the  $\beta$  estimated and modeled at month  $i$ , respectively. In this study, we use the validated *GPP residual method* to estimate  $\beta$  in tree-dominated ecosystems based on measured meteorological data and the CO<sub>2</sub> atmospheric inversion product.

2.4  $\beta$  and  $\gamma$  determined from the TRENDY v9 ensemble

We use simulations from twelve process-based global dynamic vegetation models (DGVM) within the TRENDY projects (Sitch et al., 2015; Le Quéré et al., 2018) to derive the modeled  $\beta$  and  $\gamma$ . We use four simulations (called S0, S1, S2, S3 in the TRENDY v9 protocol; see Table S3) with and without land use and land cover changes (LULCCs) under both transient (historically observed) and pre-industrial (constant) environmental conditions. CO<sub>2</sub> effect on GPP modeled by TRENDY ( $\beta_{S1-S0}$ ) is calculated as the difference between output from S1 and S0, to avoid the effect from climate recycling. To test the robustness of the *GPP residual method* and the potential LULCCs effect, we apply the same statistical method (*i.e.*, the *GPP residual method*) on simulations in S2 and S3, respectively. We derive  $\gamma$  from S2 simulations by calculating the sensitivity of GPP in S2 to temperature. We select grid-cells containing the 38 considered eddy covariance sites in all models.  $\beta$  and  $\gamma$  is calculated for the same site-months as data analyzed in EC records.

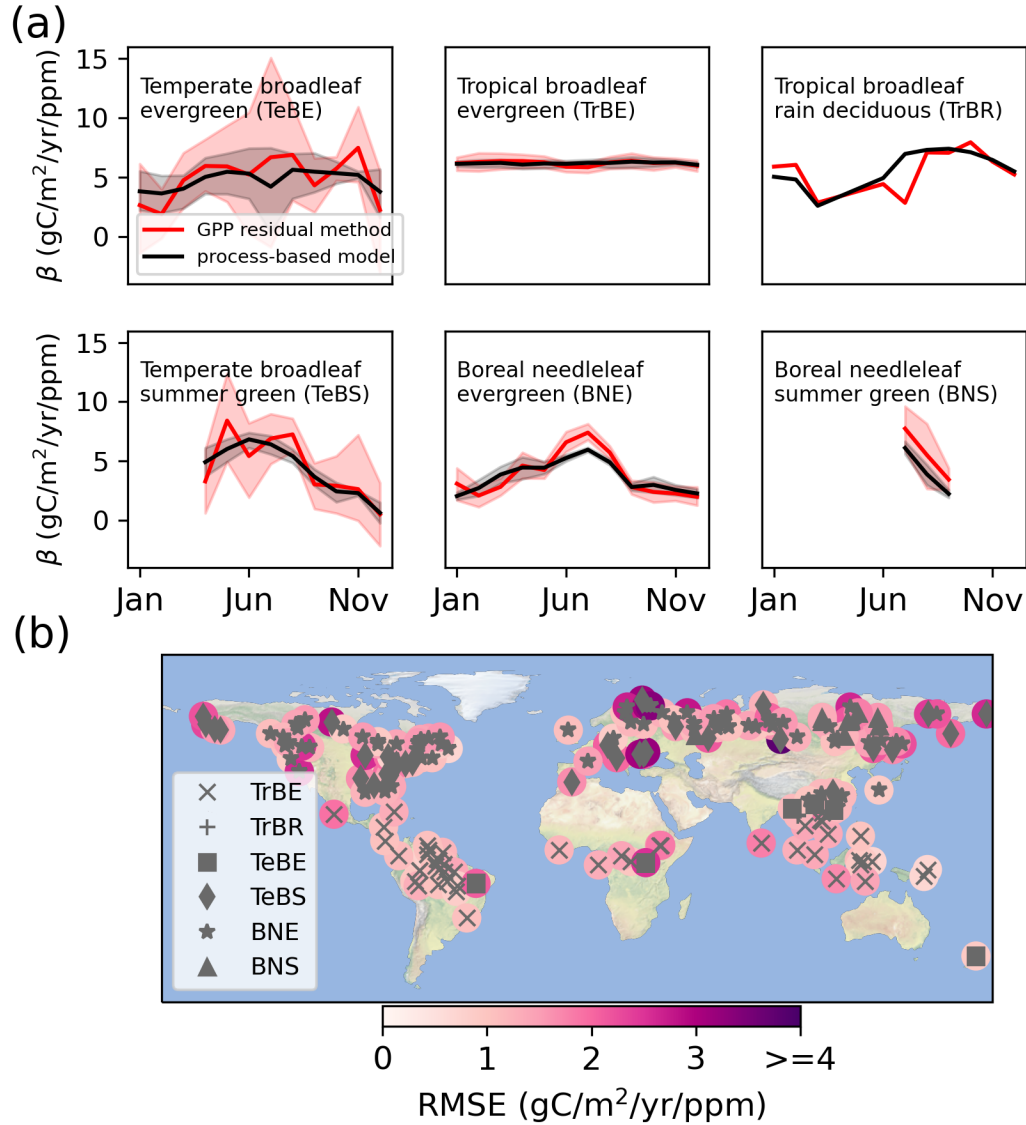
### 3 Results and Discussion

#### 3.1 Evaluating the *GPP residual method* with a land surface model

We develop and test the *GPP residual method* with QUINCY model simulations (Methods) from which we cannot only infer  $\beta$  with our method, but also directly compare with the modeled  $\beta$ . Overall, we find that our method can satisfactorily estimate  $\beta$ , and can capture the seasonal variations of  $\beta$  across biomes. We find  $\beta$  in tropical forest is overall well reproduced by our statistical method, supported by a mean root-mean-square error (RMSE) of 1.2 gC/m<sup>2</sup>/year/ppm. However, the performance in the cold northern high latitude regions, where part of the boreal

285 needle leaf forests and temperate forests are located, is slightly diminished, with a mean RMSE  
286 of 1.5 gC/m<sup>2</sup>/year/ppm (Figure 2a, b). The *GPP residual method* with implemented random  
287 forests exhibits better estimation in temperate broadleaf summergreen trees (TeBS), compared  
288 with both multivariate regression methods that estimate negative  $\beta$  in TeBS during the beginning  
289 and the late growing season (Figure S1, S2). Additionally, we find a slight overestimation  
290 accompanied with higher RMSE, during summer months in boreal needleleaf evergreen (BNE)  
291 forested sites and boreal needleleaf summergreen (BNS) forested sites. The discrepancy between  
292 estimated  $\beta$  and modeled  $\beta$  can be attributed to the limitations associated with constructing a  
293 statistical model to estimate the sensitivity of GPP to climate variables relying on interannual  
294 variabilities. This means this statistical model does not account for vegetation acclimation on  
295 climatic variability in the long-term, such as phenological changes, which cannot be learned  
296 from interannual variabilities. Thus, the statistical method exhibits decreased accuracy,  
297 particularly in ecosystems where seasonality exerts strong control.

298 In addition to the limitation of capturing vegetation phenology, we individually consider the  
299 effect of rising CO<sub>2</sub> and the effect of changing climatic conditions. Thus, the synergetic effect of  
300 rising CO<sub>2</sub> and temperature (Drake et al., 1997) is not considered in our approach, where *e.g.*,  
301 *increasing CO<sub>2</sub> can modify plants' response to temperature*. This simplification could result in  
302 the overestimation of the CO<sub>2</sub> fertilization effect on GPP. On the other hand, the anomalies  
303 associated with extreme events can be theoretically reproduced by the statistical method.  
304 However, given that only a few instances of extreme events are in the training dataset, we  
305 acknowledge that the non-linear relationship between climate and GPP during extreme  
306 conditions can induce errors in the estimation of  $\beta$ . Overall, we find encouraging consistency  
307 between the  $\beta$  estimated by the *GPP residual method* and  $\beta$  modeled by QUINCY.



**Figure 2.** Validation of the *GPP residual method* with QUINCY model simulations. (a) Seasonal variation of  $\beta$  across vegetation types estimated by the *GPP residual method* with a random forest model in red and QUINCY in black (166 sites with 1220 site-months). The red and black shaded area depicts one standard deviation around the mean value of  $\beta$  across multiple site-months (solid lines). (b) The map shows the root mean square error (RMSE) between estimated  $\beta$  and modeled  $\beta$  in the growing season for each site in the QUINCY model. Brighter color indicates lower bias and thus a better performance of the *GPP residual method*.

### 3.2 CO<sub>2</sub> fertilization effect in forested ecosystems inferred from eddy covariance records

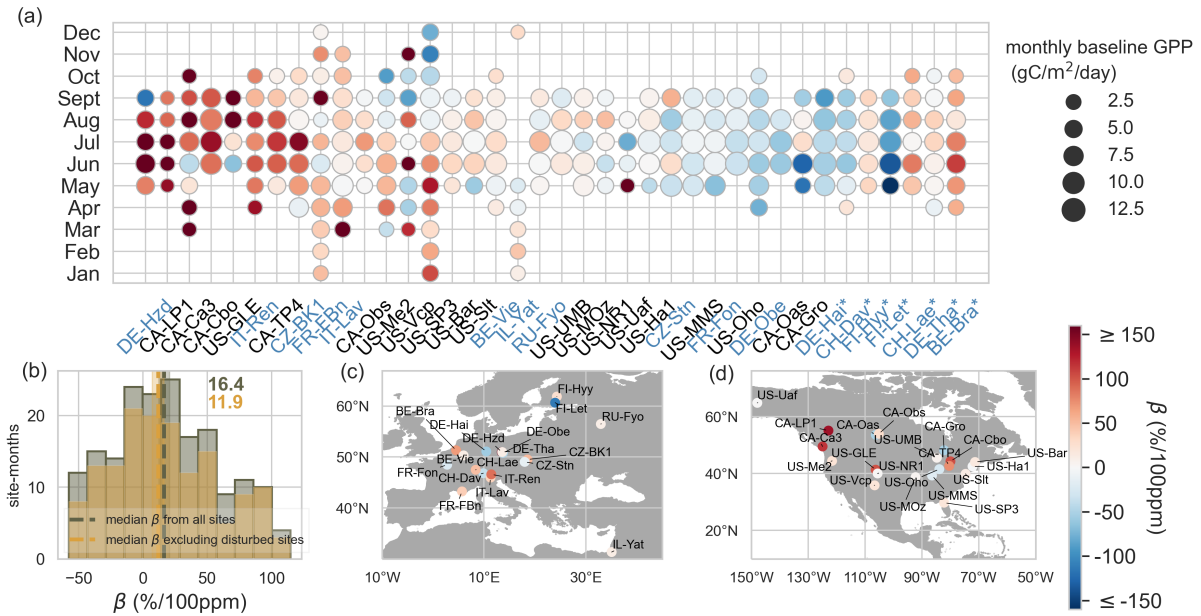
Using the *GPP residual method*, we estimate the strength of the CO<sub>2</sub> fertilization effect on

photosynthetic carbon uptake as recorded in eddy covariance (EC) time series at 38 forested sites (Figure 3). We assess the sensitivity of GPP to CO<sub>2</sub>, denoted  $\beta$  (Methods), separately for each individual month across the years of the time series to account for seasonal variations. The median  $\beta$  value across all sites and months is  $16.4 \pm 4.1$  % increase in GPP per 100 ppm rise in atmospheric CO<sub>2</sub>. While  $\beta$  displays considerable variability across sites and months, positive  $\beta$  values are consistently observed in 61 % of sites for at least two months in the record. The strongest enhancement of  $\beta$  occurs during the boreal summer months, although a selection of sites exhibits stronger effects in spring (*e.g.*, *CA-Ca3*, *IT-Ren*) or autumn (*e.g.*, *US-GLE*, *FR-FBn*).

Among the analyzed sites, the top seven sites listed in Figure 3a (*DE-Hzd*, *CA-LP1*, *CA-Ca3*, *CA-Cbo*, *US-GLE*, *IT-Ren*, *CA-TP4*) exhibit the most pronounced GPP enhancement, with their site-specific annual mean  $\beta$  values surpassing the top 20% of all the sites. In addition to the median  $\beta$  value across all sites and months, we aggregate monthly  $\beta$  based on the monthly baseline GPP, to represent the mean  $\beta$  across selected sites. The aggregated mean  $\beta$  is 14.9 % per 100 ppm, indicating the median  $\beta$  is representative and not biased by the site selection, considering the variation in GPP across sites.

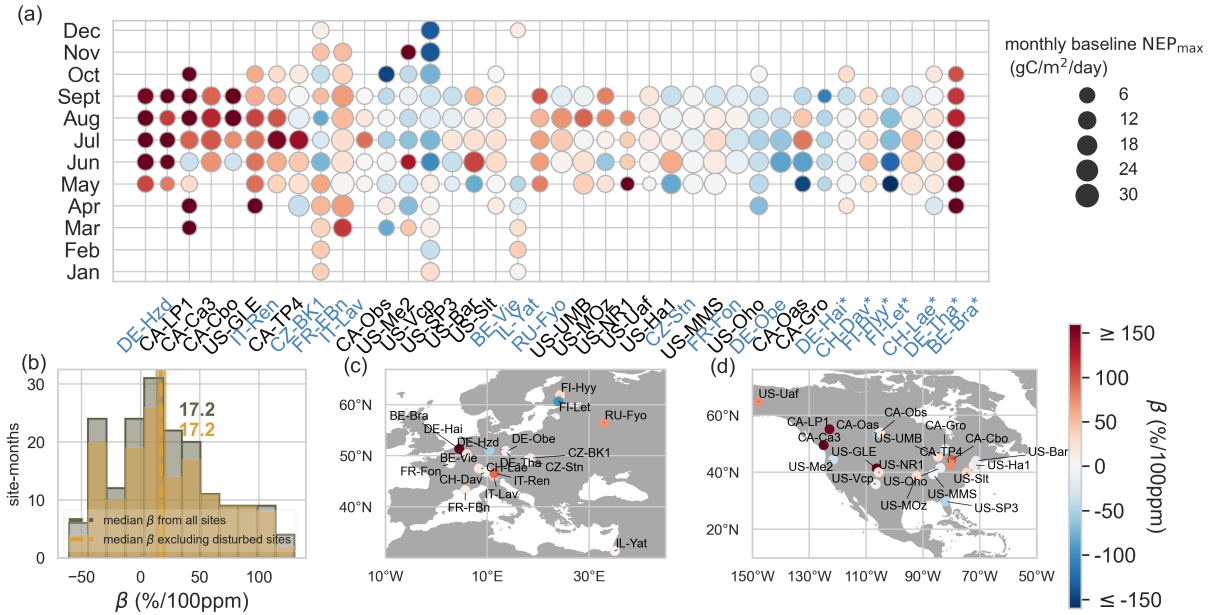
The presented approach lacks the ability to isolate additional hidden long-term effects stemming from human activities, and these effects may manifest in an over- or under-estimation of  $\beta$  derived at individual sites. Notably, certain sites (*e.g.*, *DE-Obe*, *CH-Dav*, *FI-Let*, *DE-Hai*) exhibit negative  $\beta$  values consistently throughout the growing months. The negative  $\beta$  identified at the “CH-Dav” site may be associated with a disturbance event, specifically a harvest conducted in the year 2006. Similarly, the thinning activity at the “FI-Let” site in 2016 induced a declining of GPP trend, leading to a negative  $\beta$  estimate throughout the year. Forest in the “DE-Hai” site is recovering from a severe drought. Conversely, other drivers such as nitrogen deposition at nitrogen-limited sites (De Vries et al., 2006; Sutton et al., 2008; de Vries et al., 2014) or forests undergoing succession (Pugh et al., 2019) can induce a long-term increase of GPP, potentially resulting in an overestimation of  $\beta$ . However, these rather site-specific trends in opposing directions may offset each other within a well-distributed and sufficiently large sample size, enabling the median  $\beta$  across sites to predominantly capture the widespread CO<sub>2</sub> fertilization effect. This notion is supported by excluding known disturbance sites (*e.g.*, *forest*

thinning) from the analysis, resulting in a median  $\beta$  ( $11.9 \pm 4.1$  % per 100 ppm) that does not significantly differ from the  $\beta$  estimated using all sites (Figure 3b).



**Figure 3.** Estimation of  $\beta$  from eddy covariance data using the *GPP residual method* with a random forest model. (a) Plot showing the estimated  $\beta$  for each eddy covariance site across months in the growing season. The size of circles represents the magnitude of monthly baseline GPP. Sites are shown in descending order of the annual mean  $\beta$  (Methods). Site-codes marked by a star are presented separately at the end of the list, indicating that disturbances have been recorded at those specific sites. Site-codes shown in blue and black color locate in Europe (c) and North America (d). (b) The histogram of  $\beta$  values shown in panel (a). The grey (yellow) vertical dashed lines denote the median  $\beta$  (Methods) estimated from all sites and months (excluding the disturbed sites). The grey (yellow) shaded area indicates the bootstrap estimates for the uncertainty of median  $\beta$  from all sites (excluding the disturbed sites). Maps (c) and (d) display the annual mean  $\beta$  values at each site.

The daily maximum net ecosystem production ( $NEP_{\max}$ ) provides insights into the peak photosynthetic activity of the ecosystem during optimal conditions within a day. It is valuable for understanding the ecosystem's contribution to carbon sequestration. In addition to GPP, we identify the  $CO_2$  fertilization effect on  $NEP_{\max}$  as  $17.2 \pm 3.6$  % per 100 ppm (Figure 4). The temporal and spatial variation of  $\beta$  in  $NEP_{\max}$  is consistent with  $\beta$  in GPP (Figure 3), adding additional observational evidence of the  $CO_2$  fertilization effect for better understanding of the global carbon cycle dynamics.



**Figure 4.** Estimation of  $\beta$  from eddy covariance data using the *GPP residual method* with a random forest model. Analogous to Figure 3 but  $\beta$  is estimated for  $NEP_{max}$ .

Next, we assess the robustness of our findings by testing multiple regression methods in estimating the GPP sensitivity to climatic changes, and evaluate their statistical performance. Testing a multivariate linear regression instead of a random forest regressor, we find that the median  $\beta$  yields a slightly different estimate of  $15.1 \pm 5.0$  % per 100 ppm (Figure S3). If we however apply the multivariate regression model without accounting for confounding drivers of rising CO<sub>2</sub> and climatic changes (Methods), the median  $\beta$  is notably lower and amounts to  $11.4 \pm 5.6$  % per 100 ppm (Figure. S4). We utilize the “Out-of-Bag” (OOB) score to estimate the performance of the random forest regressor on unseen data without the need for a separate validation set (Methods). Although there are instances of relatively low OOB score at certain sites and months, no clear relationship emerges between estimated  $\beta$  values and model performances (Figure S5).

### 3.3 Exploring the spatial variation of the CO<sub>2</sub> fertilization effect

We further explore the spatial variability in estimated  $\beta$ . Thereby, we assess the roles of plant functional types (PFTs), forest age, temperature, VPD and maximum leaf area index (LAI). Past studies have found stronger stomatal responsiveness to changes in CO<sub>2</sub> in deciduous trees versus conifers (Saxe et al., 1998; Medlyn et al., 2001; Brodribb et al., 2009; Klein and Ramon, 2019),

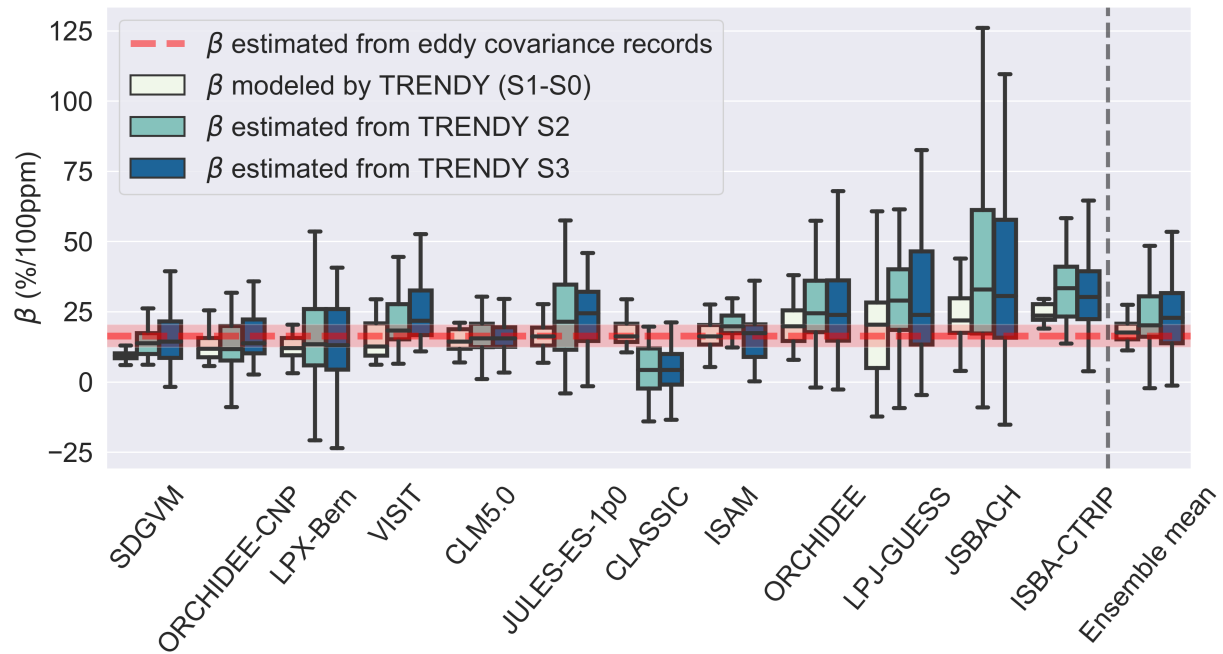


although variability exists when assessing their different responses in photosynthesis (Saxe et al., 1998). Overall, we find a greater enhancement in GPP in evergreen needle-leaved forest (ENF) in response to increasing CO<sub>2</sub> (Figure S6), compared with deciduous broadleaf forest (DBF). The difference in GPP responses to increasing CO<sub>2</sub> across PFTs may vary with scales, or complex environmental conditions (*e.g.*, *under stress or not*). Future work may focus on this difference in more detail. An open question is whether mature forests, which may be approaching a quasi-equilibrium state, are responding to CO<sub>2</sub> and climate in the same fashion as younger stands (Kira and Shidei, 1967; Odum, 1969; Luyssaert et al., 2008). We find no significant relationship between  $\beta$  and forest age, but we show a tendency of GPP enhancement to decline towards older stands (Figure S7a). Theoretically, the enhancement in GPP would relate to differences in growing season temperature and VPD, with greater enhancement at warmer growth temperatures (suppression of photorespiration; Baig et al., 2015). However, we find no significant but slight negative trend in the relationship between  $\beta$  and temperature as well as VPD (Figure S7b, c). This trend may be attributed to the combined impact of temperature and VPD. We find a positive tendency of  $\beta$  with increasing site maximum LAI (Figure S7d), which could be an interaction between rising CO<sub>2</sub> and exponential growth phase (*i.e.*, *regrowth*).

### 3.4 Comparing the CO<sub>2</sub> fertilization effect inferred from eddy covariance sites and TRENDY ensemble

We compare our EC based  $\beta$  estimates with an ensemble of twelve process-based global dynamic vegetation models (DGVM) following the TRENDY simulation protocol (Sitch et al., 2015; Le Quéré et al., 2018). The TRENDY ensemble consists of four experiments (Table S3): the pre-industrial control run (S0), the run considering only CO<sub>2</sub> changes (S1), the run considering CO<sub>2</sub> and climate change forcings (S2), and the latter with additional prescribed land-use and land-cover changes (S3). To conduct the comparison, we extract model time series from the individual grid-cells containing the 38 considered eddy-covariance sites. The modeled CO<sub>2</sub> fertilization effect inferred by calculating the difference  $\beta_{S1-S0}$ , exhibits a large spread among the TRENDY models. The median  $\beta_{S1-S0}$  across the grid-cells and models is 17.7 % per 100 ppm, which is remarkably close to the median  $\beta$  obtained through the *GPP residual method* using EC records. Seven models (VISIT, CLM5.0, JULES-ES-1p0, CLASSIC, ISAM, ORCHIDEE and LPJ-GUESS) fall within the bootstrapped uncertainty range of the median  $\beta$  estimated from

observations (Figure 5; Methods). We acknowledge the limitation of this comparison, as it involves contrasting site-level estimates with grid-level results, which is influenced by the heterogeneity within each grid-cell. Nevertheless, we argue that the median values across the sites and grid-cells provide a more aggregated perspective that helps mitigate the influence of sub-grid heterogeneity.



**Figure 5.** Comparing  $\beta$  estimated from eddy covariance data and the TRENDY model ensemble. The medians and interquartile ranges of  $\beta$  are shown for each model and for the ensemble mean, as horizontal lines within the boxes, and the upper and bottom lines of the box, respectively. Each box includes grid-cells containing the 38 considered eddy covariance sites. Box plots for individual models are in an ascending order based to the median  $\beta$ . The dotted red line represents the median  $\beta$  derived from eddy covariance records (as shown also in Figure 3), with the uncertainty showing in shaded area.

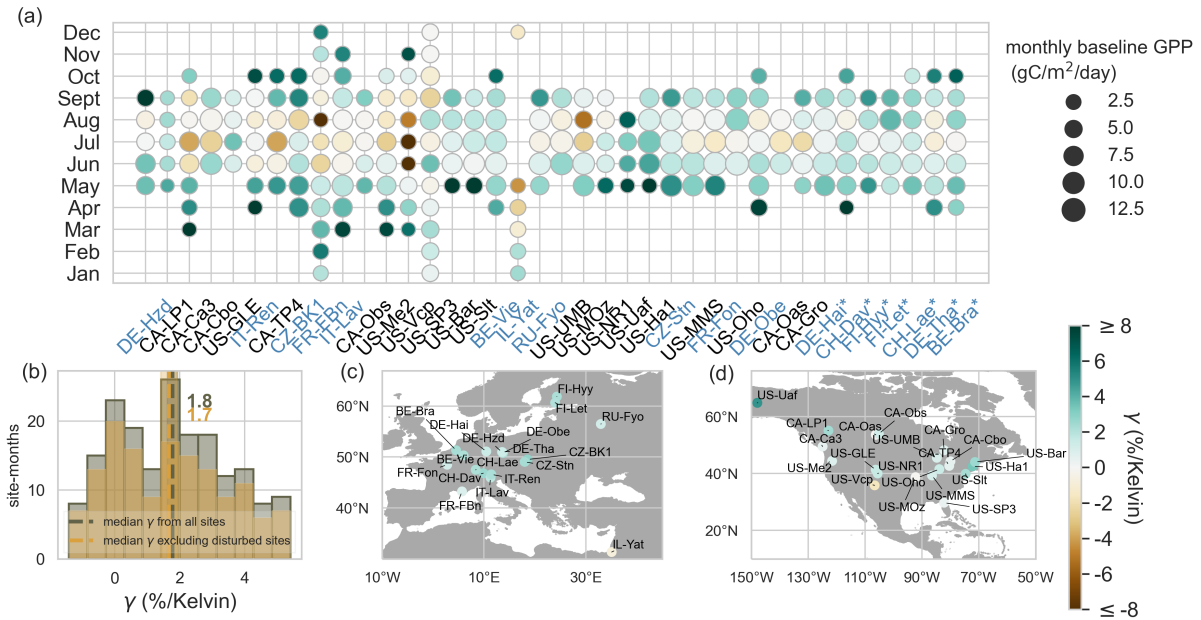
Other factors, such as nitrogen deposition, disturbances, and particularly land management, can influence ecosystem productivity as recorded in EC data. These factors potentially influence the estimation of  $\beta$  using the *GPP residual method*. To assess the effect of land-use and land-cover changes (LULCCs) on  $\beta$  estimation, we compare  $\beta$  derived from the TRENDY S2 and S3 simulations using the *GPP residual method*. With the exception of the LPJ-GUESS and JULES-ES-1p0 models, the ensemble mean of  $\beta_{S2}$  closely aligns with  $\beta_{S3}$  (Figure 5), suggesting that the neglected effects of LULCCs do not substantially affect the  $\beta$  estimation. Furthermore, in line

with the validation using QUINCY simulations, the *GPP residual method* tends to slightly overestimate  $\beta$  when comparing  $\beta_{S2}$  with  $\beta_{S1-S0}$  derived from TRENDY. This further emphasizes that the method cannot account for long-term vegetation acclimation and phenological changes; however, these effects are minor within the considered time period.

### 3.5 Influence of climatic changes on productivity throughout the season

Conventionally, the  $\gamma$  factor is defined as the sensitivity of land carbon storage to climate variations using temperature change as the proxy (Friedlingstein et al., 2006; Gregory et al., 2009; Arora et al., 2020). Analogously, we define  $\gamma$  as the relative change in the climate-driven GPP component over temperature change, which is already obtained in the *GPP residual method* (Methods). The median  $\gamma$  in GPP estimated from the EC dataset is  $1.8 \pm 0.2$  % per Kelvin (Figure 5). Comparing this to the sensitivity of GPP to  $\text{CO}_2$ , assuming a 100 pm increase in atmospheric  $\text{CO}_2$  concentration is roughly equivalent to 1 Kelvin temperature increase in the historical period, we find that  $\gamma$  is considerably lower than  $\beta$ , in line with previous studies (Fernández-Martínez et al., 2017; Chen et al., 2022). The median  $\gamma$  in  $\text{NEP}_{\text{max}}$  ( $-0.2 \pm 0.2$  % per Kelvin) is much lower than  $\gamma$  in GPP, reflecting a negative response of  $\text{NEP}_{\text{max}}$  to temperature variations, particularly at the peak of the growing season (Figure S8). The median  $\gamma$  estimated from the TRENDY ensemble (S2 simulations) is  $0.8 \pm 0.2$  % per Kelvin (Figure S9). Also,  $\gamma$  exhibits a large spread among models compared to  $\gamma$  from EC, suggesting a more pronounced uncertainty in the process representation in estimating ecosystem responses to climate changes among the land-surface models (Figure S10). A clear seasonality of  $\gamma$  emerges in both observations and models (Figure 5, Figure S8, Figure S9). While  $\gamma$  is higher at the beginning and the end of the growing season for most of the sites, most sites show negative  $\gamma$  in at least one month of the growing season (26 out of 38 in the EC estimated GPP; 36 out of 38 in the EC measured  $\text{NEP}_{\text{max}}$ ; 33 out of 38 in the TRENDY ensemble mean). Our results indicate that warming may have a positive effect on vegetation productivity at colder conditions and a potential negative effect in warm climate. A high temperature is usually accompanied by a high VPD, which limits the stomatal conductance and evapotranspiration (Park Williams et al., 2013; Novick et al., 2016). On the other hand, plant productivity response to temperature is associated with water availability; a positive effect of temperature can occur when water is not limiting the ecosystem functioning (Fernández-Martínez et al., 2019). Additionally, the negative climate-

carbon feedback may contribute to, and potentially partly mask or enhance the CO<sub>2</sub> fertilization effect.



**Figure 6.** Estimation of  $\gamma$  from eddy covariance dataset using the *GPP residual method* with a random forest model. (a) Plot showing the estimated  $\gamma$  for each eddy covariance site across months in the growing season. The size of circles represents the magnitude of monthly baseline GPP. Sites are shown in descending order of the annual mean  $\gamma$  (Methods). Site-codes marked by a star are presented separately at the end of the list, indicating that disturbances have been recorded at those specific sites. Site-codes shown in blue and black color locate in Europe (c) and North America (d). (b) The histogram of  $\gamma$  values shown in panel (a). The grey (yellow) vertical dashed lines denote the median  $\gamma$  (Methods) estimated from all sites and months (excluding the disturbed sites). The grey (yellow) shaded area indicates the bootstrap estimates for the uncertainty of median  $\gamma$  from all sites (excluding the disturbed sites). Maps (c) and (d) display the annual mean  $\gamma$  values at each site.

Overall, we recognize the inherent limitations in EC-based data acquisition, the assumptions of the *GPP residual method*, and the potential influence of other long-term factors such as human activities, which can introduce biases in the estimation of  $\beta$  and  $\gamma$  from observations. However, despite these challenges, the  $\beta$  and  $\gamma$  values estimated from the EC records align with the  $\beta$  and  $\gamma$  values simulated by the TRENDY model, both in terms of the median values across multiple sites and the seasonality of  $\gamma$ . Furthermore, utilizing the TRENDY simulations, we demonstrate that the discrepancy in  $\beta$  estimation, influenced by land use and land cover changes, remains within an acceptable range.

## 490    **4 Conclusions**

491    Our study isolates a robust, multi-decadal enhancement in vegetation productivity ( $\beta_{GPP} = 16.4$   
 492     $\pm 4 \%$  per 100 ppm,  $\beta_{NEP_{max}} = 17.2 \pm 4 \%$  per 100 ppm) across Northern Hemisphere forests in  
 493    response to the rising atmospheric CO<sub>2</sub> concentration. We further diagnose the median value of  
 494    GPP sensitivity to temperature ( $\gamma$ ) of  $1.8 \pm 0.2 \%$  per Kelvin, and find evidence of a negative  
 495    effect of temperature on photosynthesis at the peak of the growing season. Assuming a 100 pm  
 496    increase in CO<sub>2</sub> concentration is equivalent to 1 Kelvin temperature increase, the negative  
 497    temperature effect potentially masks the positive increasing CO<sub>2</sub> effect on GPP. While the  
 498    TRENDY ensemble captures the median  $\beta$  and  $\gamma$  inferred from eddy covariance records, there is  
 499    a notable variation in  $\beta$  and  $\gamma$  exhibited between individual models. To improve the  
 500    representation of the temperature effect on GPP in TRENDY models, further studies should  
 501    focus on reducing uncertainties associated with physiological processes and incorporating multi-  
 502    model constraints. This study paves the way for future investigations into long-term drivers of  
 503    change and ecosystem functioning. We anticipate that our approach could be readily applied to  
 504    other ecosystems (*e.g.*, *drylands*), other datasets (long-term satellite records of change, *i.e.*,  
 505    *vegetation greenness, etc.*), and other variables that describe ecosystem function (*e.g.*,  
 506    *evapotranspiration*). Together with a wide range of evidence of global vegetation productivity in  
 507    response to rising atmospheric CO<sub>2</sub> concentration and climate change, a better estimation of the  
 508    remaining carbon budget to achieve the climate goals in the Paris Agreement is possible.

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## Open Research

The observational data that support the findings of this study are available in the ICOS warm winter 2020 product (<https://www.icos-cp.eu/data-products/2G60-ZHAK>) and Ameriflux (<https://ameriflux.lbl.gov/>).

The codes are available online: <https://doi.org/10.5281/zenodo.10125474>.

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